PROBABILITY THEORY LECTURE 1

Per Sidén

Division of Statistics and Machine Learning Dept. of Computer and Information Science Linköping University

PER SIDÉN (STATISTICS, LIU)

OVERVIEW LECTURE 1

- Course outline
- Introduction and a recap of some background
- Functions of random variables
- Multivariate random variables

COURSE OUTLINE

- ▶ 6 Lectures: theory interleaved with illustrative solved examples.
- ▶ 6 Seminars: problem solving sessions + open discussions.
- ▶ 1 Recap session: Recap of the course.

COURSE LITERATURE

- Gut, A. An intermediate course in probability. 2nd ed. Springer-Verlag, New York, 2009. ISBN 978-1-4419-0161-3
- Chapter 1: Multivariate random variables
- Chapter 2: Conditioning
- Chapter 3: Transforms
- Chapter 4: Order statistics
- Chapter 5: The multivariate normal distribution
- Chapter 6: Convergence

EXAMINATION

The examination consists of a written exam with max score 20 points and grade limits:

A: 19p, B: 17p, C: 14p, D: 12p, E: 10p.

- You are allowed to bring a pocket calculator to the exam, but no books or notes.
- The following will be distributed with the exam:
 - Table with common formulas and moment generating functions (available on the course homepage).
 - ► Table of integrals (available on the course homepage).
 - ► Table with distributions from Appendix B in the course book.
- Active participation in the seminars gives bonus points to the exam. A student who earns the bonus points will add 2 points to the exam result in order to reach grade E, D or C, 1 point in order to reach grade B, but no points in order to reach grade A. Required exam results for a student who earned the bonus points for respective grade: A: 19p, B: 16p, C: 12p, D: 10p, E: 8p.

PER SIDÉN (STATISTICS, LIU)

BONUS POINTS

- To earn the bonus points a student must be present and active in at least 5 of the 6 seminars, so maximally one seminar can be missed regardless of reasons.
- Active participation means that the student has made an attempt to solve every exercise indicated in the timetable before respective seminar and is able to present his/her solutions on the board during the seminar. Active participation also means that the student gives help and comments to the classmates' presented solutions.
- In the seminars, for each exercise a student will be randomly selected to present his/her solution (without replacement).
- Exercises marked with * are a bit harder and it is ok if you are not able to solve these.

COURSE HOMEPAGE

https://www.ida.liu.se/~732A63/ (select english)

RANDOM VARIABLES

- The sample space Ω = {ω₁, ω₂, ...} of an experiment is the most basic representation of a problem's randomness (uncertainty).
- ▶ More convenient to work with real-valued measurements.
- A random variable X is a real-valued function from a sample space: $X = f(\omega)$, where $f : \Omega \to \mathbb{R}$.
- A multivariate random vector: $\mathbf{X} = f(\omega)$ such that $f : \Omega \to \mathbb{R}^n$.

Examples:

• Roll a die: $\Omega = \{1, 2, 3, 4, 5, 6\}.$

$$X(\omega) = \begin{cases} 0 & \text{if } \omega = 1,2 \text{ or } 3\\ 1 & \text{if } \omega = 4,5 \text{ or } 6 \end{cases}$$

▶ Roll two fair dice. $X(\omega)$ =sum of the two dice.

PER SIDÉN (STATISTICS, LIU)

SAMPLE SPACE OF TWO DICE EXAMPLE



PER SIDÉN (STATISTICS, LIU)

THE DISTRIBUTION OF A RANDOM VARIABLE

- The probabilities of events on the sample space Ω imply a probability distribution for a random variable X(ω) on Ω.
- The probability distribution of X is given by

$$\Pr(X \in C) = \Pr(\{\omega : X(\omega) \in C\}),\$$

where $\{\omega : X(\omega) \in C\}$ is the event (in Ω) consisting of all outcomes ω that gives a value of *X* in *C*.

- A random variable is discrete if it can take only a finite or a countable number of different values x₁, x₂,
- **Continuous** random variables can take every value in an interval.

DISCRETE RANDOM VARIABLE

 The probability function (p.f), is the function

$$p(x) = \Pr(X = x)$$



UNIFORM, BERNOULLI AND POISSON

• **Uniform discrete** distribution. $X \in \{a, a + 1, ..., b\}$.

$$p(x) = \begin{cases} \frac{1}{b-a+1} & \text{for } x = a, a+1..., b\\ 0 & \text{otherwise} \end{cases}$$

- ▶ Bernoulli distribution. $X \in \{0, 1\}$. Pr(X = 0) = 1 p and Pr(X = 1) = p.
- ▶ Poisson distribution: $X \in \{0, 1, 2, ...\}$

$$p(x) = \frac{\exp(-\lambda) \cdot \lambda^x}{x!} \quad \text{for } x = 0, 1, 2, \dots$$

PER SIDÉN (STATISTICS, LIU)

THE BINOMIAL DISTRIBUTION

Binomial distribution. Sum of n independent Bernoulli variables X₁, X₂, ..., X_n with the same success probability p.

$$X = X_1 + X_2 + \dots + X_n$$
$$X \sim Bin(n, p)$$

Probability function for a Bin(n, p) variable:

$$P(X = x) = {n \choose x} p^{x} (1 - p)^{n-x}$$
, for $x = 0, 1, ..., n$.

► The binomial coefficient ⁿ_x is the number of binary sequences of length n that sum exactly to x.

PER SIDÉN (STATISTICS, LIU)

PROBABILITY DENSITY FUNCTIONS

- > Continuous random variables can assume every value in an interval.
- Probability density function (pdf) f(x)

•
$$\Pr(a \le X \le b) = \int_a^b f(x) dx$$



•
$$f(x) \ge 0$$
 for all x

•
$$\int_{-\infty}^{\infty} f(x) dx = 1$$

A pdf is like a histogram with tiny bin widths. Integral replaces sums.

Continuous distributions assign probability zero to individual values, but

$$\Pr\left(\mathbf{a} - \frac{\epsilon}{2} \le X \le \mathbf{a} + \frac{\epsilon}{2}\right) \approx \epsilon \cdot f(\mathbf{a}).$$

PER SIDÉN (STATISTICS, LIU)

DENSITIES - SOME EXAMPLES

► The **uniform** distribution

$$f(x) = egin{cases} rac{1}{b-a} & ext{ for } a \leq x \leq b \ 0 & ext{ otherwise.} \end{cases}$$

► The triangle or linear pdf

$$f(x) = \begin{cases} \frac{2}{a^2}x & \text{ for } 0 < x < a \\ 0 & \text{ otherwise} \end{cases}$$

▶ The normal, or Gaussian, distribution

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} \left(x - \mu\right)^2\right)$$

PER SIDÉN (STATISTICS, LIU)

EXPECTED VALUES, MOMENTS

The expected value of X is

$$E(X) = \begin{cases} \sum_{k=i}^{\infty} x_k \cdot p(x_k) & , X \text{ discrete} \\ \int_{-\infty}^{\infty} x \cdot f(x) & , X \text{ continuous} \end{cases}$$

- Example: E(X) when $X \sim Uniform(a, b)$
- The *n*th moment is defined as $E(X^n)$
- The variance of X is $Var(X) = E(X EX)^2 = E(X^2) (EX)^2$

THE CUMULATIVE DISTRIBUTION FUNCTION

► The (cumulative) distribution function (cdf) F(·) of a random variable X is the function

$$F(x) = \Pr(X \le x) ext{ for } -\infty \le x \le \infty$$

- Same definition for discrete and continuous variables.
- ► The cdf is non-decreasing

If
$$x_1 \leq x_2$$
 then $F(x_1) \leq F(x_2)$

- Limits at $\pm \infty$: $\lim_{x \to -\infty} F(x) = 0$ and $\lim_{x \to \infty} F(x) = 1$.
- For continuous variables: relation between pdf and cdf

$$F(x) = \int_{-\infty}^{x} f(t) dt$$

and conversely

$$\frac{dF(x)}{dx} = f(x)$$

PER SIDÉN (STATISTICS, LIU)

FUNCTIONS OF RANDOM VARIABLES

- ► Quite common situation: You know the distribution of X, but need the distribution of Y = g(X), where g(·) is some function.
- Example 1: $Y = a + b \cdot X$, where a and b are constants.
- Example 2: Y = 1/X
- Example 3: $Y = \ln(X)$.
- Example 4: $Y = \log \frac{X}{1-X}$
- Y = g(X), where X is discrete.
- $p_X(x)$ is p.f. for X. $p_Y(y)$ is p.f. for Y:

$$p_{Y}(y) = \Pr(Y = y) = \Pr[g(X) = y] = \sum_{x:g(x)=y} p_{X}(x)$$

PER SIDÉN (STATISTICS, LIU)

FUNCTION OF A CONTINUOUS RANDOM VARIABLE

▶ Suppose that X is continuous with support (a, b). Then

$$F_{Y}(y) = \Pr(Y \le y) = \Pr[g(X) \le y] = \int_{x:g(x) \le y} f_{X}(x) dx$$

▶ Let g(X) be monotonically *increasing* with inverse X = h(Y). Then $F_Y(y) = Pr(Y \le y) = Pr(g(X) \le y) = Pr(X \le h(y)) = F_X(h(y))$ and $P_Y(y) = P_Y(Y \le y) = P_Y(y) = P_Y(y) = P_Y(y)$

$$f_Y(y) = f_X(h(y)) \cdot \frac{\partial h(y)}{\partial y}$$

For general monotonic transformation Y = g(X) we have

$$f_Y(y) = f_X[h(y)] \left| \frac{\partial h(y)}{y} \right|$$
 for $\alpha < y < \beta$

where (α, β) is the mapped interval from (a, b).

PER SIDÉN (STATISTICS, LIU)

EXAMPLES: FUNCTIONS OF A RANDOM VARIABLE

• Example 1. $Y = a \cdot X + b$.

$$f_{Y}(y) = \frac{1}{|a|} f_{X}\left(\frac{y-b}{a}\right)$$

► Example 2: log-normal. $X \sim N(\mu, \sigma^2)$. $Y = g(X) = \exp(X)$. $X = h(Y) = \ln Y$.

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{1}{2\sigma^2} \left(\ln y - \mu\right)^2\right) \cdot \frac{1}{y} \text{ for } y > 0.$$

• Example 3. $X \sim LogN(\mu, \sigma^2)$. $Y = a \cdot X$, where a > 0. X = h(Y) = Y/a.

$$f_{Y}(y) = \frac{1}{y/a} \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{1}{2\sigma^{2}} \left(\ln\frac{y}{a} - \mu\right)^{2}\right) \frac{1}{a}$$
$$= \frac{1}{y} \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{1}{2\sigma^{2}} \left(\ln y - \mu - \ln a\right)^{2}\right)$$

which means that $Y \sim LogN(\mu + \ln a, \sigma^2)$.

PER SIDÉN (STATISTICS, LIU)

EXAMPLES: FUNCTIONS OF A RANDOM VARIABLE

Example 4.
$$X \sim LogN(\mu, \sigma^2)$$
. $Y = X^a$, where $a \neq 0$.
 $X = h(Y) = Y^{1/a}$.

$$f_{Y}(y) = \frac{1}{y^{1/a}} \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{1}{2\sigma^{2}} \left(\ln y^{1/a} - \mu\right)^{2}\right) \frac{1}{a} y^{1/a-1} \cdot \\ = \frac{1}{y} \frac{1}{\sqrt{2\pi}a\sigma} \exp\left(-\frac{1}{2a^{2}\sigma^{2}} \left(\ln y - a\mu\right)^{2}\right)$$

which means that $Y \sim LogN(a\mu, a^2\sigma^2)$.

PER SIDÉN (STATISTICS, LIU)

PROBABILITY THEORY - L1

21 / 30

BIVARIATE DISTRIBUTIONS

The joint (or bivariate) distribution of the two random variables X and Y is the collection of all probabilities of the form

 $\Pr\left[(X,Y)\in C\right]$

- Example 1:
 - X = # of visits to doctor.
 - Y = #visits to emergency.
 - C may be $\{(x, y) : x = 0 \text{ and } y \ge 1\}$.
- Example 2:
 - X = monthly percentual return to SP500 index
 - Y =monthly return to Stockholm index.
 - C may be $\{(x, y) : x < -10 \text{ and } y < -10\}$.
- Discrete random variables: joint probability function (joint p.f.)

$$f_{X,Y}(x,y) = \Pr(X = x, Y = y)$$

such that $\Pr[(X, Y) \in C] = \sum_{(x,y)\in C} f_{X,Y}(x, y)$ and $\sum_{AII} (x,y) f_{X,Y}(x, y) = 1.$

PER SIDÉN (STATISTICS, LIU)

CONTINUOUS JOINT DISTRIBUTIONS

Continuous joint distribution (joint p.d.f.)

$$\Pr[(X,Y) \in C] = \iint_C f_{X,Y}(x,y) dx dy,$$

where $f_{X,Y}(x, y) \ge 0$ is the joint density.

Univariate distributions: probability is area under density.



Bivariate distributions: probability is volume under density.



▶ Be careful about the regions of integration. Example:
C = {(x, y) : x² ≤ y ≤ 1}

PER SIDÉN (STATISTICS, LIU)

EXAMPLE

► Example

$$f_{X,Y}(x,y)=rac{3}{2}y^2$$
 for $0\leq x\leq 2$ and $0\leq y\leq 1.$



BIVARIATE NORMAL DISTRIBUTION

The most famous of them all: the bivariate normal distribution, with pdf

$$f_{X,Y}(x,y) = \frac{1}{2\pi(1-\rho^2)^{1/2}\sigma_x\sigma_y} \times \exp\left(-\frac{1}{2(1-\rho^2)}\left[\left(\frac{x-\mu_x}{\sigma_x}\right)^2 - 2\rho\left(\frac{x-\mu_x}{\sigma_x}\right)\left(\frac{y-\mu_y}{\sigma_y}\right) + \left(\frac{y-\mu_y}{\sigma_y}\right)^2\right]\right)$$

Five parameters: μ_x , μ_y , σ_x , σ_y and ρ .



PER SIDÉN (STATISTICS, LIU)

BIVARIATE C.D.F.

► Joint cumulative distribution function (joint c.d.f.):

$$F_{X,Y}(x,y) = \Pr(X \le x, Y \le y)$$

► Calculating probabilities of rectangles Pr(a < X ≤ b and c < Y ≤ d):</p>

$$F_{X,Y}(b,d) - F_{X,Y}(a,d) - F_{X,Y}(b,c) + F_{X,Y}(a,c)$$

Properties of the joint c.d.f.

• Marginal of X:
$$F_X(x) = \lim_{y \to \infty} F_{X,Y}(x, y)$$

$$F_{X,Y}(x,y) = \int_{-\infty}^{y} \int_{-\infty}^{x} f_{X,Y}(r,s) drds$$

•
$$f_{X,Y}(x,y) = \frac{\partial^2 F_{X,Y}(x,y)}{\partial x \partial y}$$

PER SIDÉN (STATISTICS, LIU)

MARGINAL DISTRIBUTIONS

Marginal p.f. of a bivariate distribution is

$$f_X(x) = \sum_{All \ y} f_{X,Y}(x, y) \text{ [Discrete case]}$$
$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x, y) dy \text{ [Continuous case]}$$

A marginal distribution for X tells you about the probability of different values of X, averaged over all possible values of Y.

PER SIDÉN (STATISTICS, LIU)

INDEPENDENT VARIABLES

> Two random variables are independent if

```
Pr(X \in A \text{ and } Y \in B) = Pr(X \in A) \cdot Pr(Y \in B)
```

for all sets of real numbers A and B (such that $\{X \in A\}$ and $\{Y \in B\}$ are events).

Two variables are independent if and only if the joint density can be factorized as

$$f_{X,Y}(x,y) = h_1(x) \cdot h_2(y)$$

- Note: this factorization must hold for all values of x and y. Watch out for non-rectangular support!
- ➤ X and Y are independent if learning something about X (e.g. X > 2) has no effect on the probabilities for different values of Y.

MULTIVARIATE DISTRIBUTIONS

- Obvious extension to more than two random variables, $X_1, X_2, ..., X_n$.
- ► Joint p.d.f.

$$f(x_1, x_2, ..., x_n)$$

Marginal distribution of x₁

$$f_1(x_1) = \int_{x_2} \cdots \int_{x_n} f(x_1, x_2, ..., x_n) dx_2 \cdots dx_n$$

Marginal distribution of x₁ and x₂

$$f_{12}(x_1, x_2) = \int_{x_3} \cdots \int_{x_n} f(x_1, x_2, ..., x_n) dx_3 \cdots dx_n$$

and so on.

PER SIDÉN (STATISTICS, LIU)

FUNCTIONS OF RANDOM VECTORS

- ▶ Let X be an *n*-dimensional continuous random variable
- Let **X** have density $f_{\mathbf{X}}(\mathbf{x})$ on support $S \subset \mathbb{R}^n$.
- Let Y = g(X), where $g : S \to T \subset \mathbb{R}^n$ is a bijection (1:1 and onto).
- Assume g and g^{-1} are continuously differentiable with Jacobian

$$\mathbf{J} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \cdots & \frac{\partial x_1}{\partial y_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \cdots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

THEOREM

("The transformation theorem") The density of Y is

$$f_{\mathbf{Y}}(\mathbf{y}) = f_{\mathbf{X}} \left[h_1(\mathbf{y}), h_2(\mathbf{y}), ..., h_n(\mathbf{y}) \right] \cdot |\mathbf{J}|$$

where $h = (h_1, h_2, ..., h_n)$ is the unique inverse of $g = (g_1, g_2, ..., g_n)$.

PER SIDÉN (STATISTICS, LIU)