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Deep learning-based estimation of time-dependent parameters in Markov models with application to SDEs - theoretical foundations

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Outline



- I. Introduction - inspiration
- II. General setting - discrete Markov processes
- III. Application 1 - multivariate regression
- IV. Application 2 - estimation of time-dependent parameters in SDEs-based models

I. Introduction - inspiration



- **PROBLEM:**
approximation of time-dependent parameters (from real data)
- **IDEA:**
approximation task → optimization problem (maximum likelihood approach)
- We have good tools, because it fits into the **deep learning framework**. We can find appropriate neural network, train it to fit our data and obtain the closest possible values.

I. Introduction - inspiration

[1] O. Dürr, B. Sick, E. Murina, *Probabilistic Deep Learning*, Manning, 2020
Chapter 4 → Building loss functions with the likelihood approach

- **CASE 1:** linear relationship between input and output
- **CASE 2:** non-linear relationship between input and output, const. variance

$$\text{loss} = \text{MSE} = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

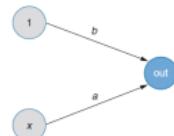


Figure 4.11. Simple linear regression as an fANN without a hidden layer. This model computes the output directly from the input as $\text{out} = a \cdot x + b$.

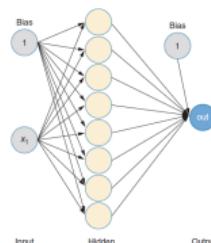


Figure 4.15. Extended linear regression. The eight neurons in the hidden layer give the features from which the output "out" is computed.

Figure: Neural network architectures for case 1 and case 2. Source: [1].

I. Introduction - inspiration

[1] O. Dürr, B. Sick, E. Murina, *Probabilistic Deep Learning*, Manning, 2020
Chapter 4 → Building loss functions with the likelihood approach

- **CASE 3:** non-linear relationship between input and output, nonconstant variance

$$\text{loss} = \text{NLL} = \sum_{i=1}^n -\log \left(\frac{1}{\sqrt{2\pi\sigma_{x_i}^2}} \right) + \frac{(\mu_{x_i} - y_i)^2}{2\sigma_{x_i}^2}$$

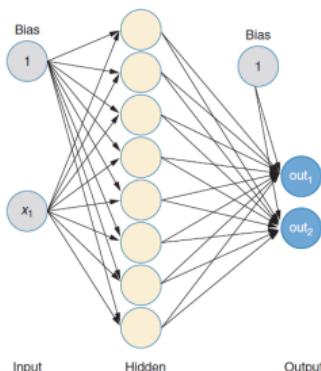


Figure 4.16 You can use an NN with two output nodes to control the parameters μ_x and σ_x of the conditional outcome distribution $N(\mu_x, \sigma_x)$ for regression tasks with nonconstant variance.

Figure: Neural network architectures for case 3. Source: [1].

II. General setting - discrete Markov processes



- We have x_1, x_2, \dots, x_n - observed realization of the sequence of random variables X_1, X_2, \dots, X_n ,
- We assume (discrete) Markov property, i.e.:

$$\mathbb{P}(X_{k+1} \in A \mid X_1, \dots, X_k) = \mathbb{P}(X_{k+1} \in A \mid X_k),$$

for all Borel sets A and $k = 0, 1, \dots, n - 1$, where

$$\mathbb{P}(X_1 \in A \mid X_0) := \mathbb{P}(X_1 \in A)$$

II. General setting - discrete Markov processes



- We assume that for all $k = 0, 1, \dots, n - 1$ there exists conditional density $f_{X_{k+1}|X_k}(\cdot | x_k)$, with $f_{X_1|X_0}(\cdot | x_0) := f_{X_1}(\cdot)$ and $f_{X_{k+1}|X_k} = f_{X_{k+1}|X_k}(\cdot | x_k, \Theta)$, $f_{X_{k+1}|X_k} : \mathbb{R}^d \times \mathbb{R}^d \times B([0, T]) \rightarrow \mathbb{R}$
- Transition densities might depend on the parameter function $\Theta : [0, T] \rightarrow \mathbb{R}^s$ in the functional way, and this dependence might be nonlinear.
- By the Markov property:

$$f_{(X_1, X_2, \dots, X_n)}(x_1, x_2, \dots, x_n, \Theta) = \prod_{k=0}^{n-1} f_{X_{k+1}|X_k}(x_{k+1} | x_k, \Theta).$$

II. General setting - discrete Markov processes



MaxLike approach:

$$\begin{aligned}\mathcal{L}(\Theta) &= -\ln f_{(X_1, X_2, \dots, X_n)}(x_1, x_2, \dots, x_n, \Theta) \\ &= -\sum_{k=0}^{n-1} \ln \left[f_{X_{k+1}|X_k}(x_{k+1}|x_k, \Theta) \right]\end{aligned}$$

Aim

$$\Theta^* = \operatorname{argmin}_{\Theta \in B([0, T])} \mathcal{L}(\Theta)$$

Problem: number of optimization parameters - in this setting the problem is infinite dimensional.

II. General setting - discrete Markov processes



NN-MaxLike approach

$$\mathcal{L}(w) = - \sum_{k=0}^{n-1} \ln \left[f_{X_{k+1}|X_k}(x_{k+1}|x_k, \Theta(w)) \right],$$

- Θ is modeled by artificial neural network
- customized loss function \mathcal{L} is the crucial part
- weights $w \in \mathbb{R}^N$ and we assume that N is strictly smaller than n

(Better!) Aim

$$w^* = \operatorname{argmin}_{w \in \mathbb{R}^N} \mathcal{L}(w)$$

Number of optimization parameters is $O(N)$, which depends on architecture of NN but is independent of n .

III. Application 1 - multivariate regression

- x_1, x_2, \dots, x_n - realizations of the \mathbb{R}^d -valued random vectors

$$X(t_k) = \mu(t_k) + \varepsilon(t_k), \quad k = 1, 2, \dots, n \quad (1)$$

- $(\varepsilon(t_k))_k$ - independent random variables $\sim N(0, \Sigma(t_k))$ with $\Sigma(t_k) \in \mathbb{R}^{d \times d}$, $\mu(t_k) \in \mathbb{R}^d$.
- The symmetric covariance matrix $\Sigma(t_k)$ - strictly positive definite for all $k = 1, 2, \dots, n$.

$$\Theta(t) = [\mu(t), \Sigma(t)], \quad \Theta(t, w) = [\mu(t, w)), \Sigma(t, w)]$$

$(X(t_k))_k$ - independent and normally distributed, hence for $x \in \mathbb{R}^d$.

$$\begin{aligned} f_{X_k|X_{k-1}}(x|x_{k-1}, \Theta) &= f_{X_k}(x|\Theta) = (2\pi)^{-d/2}(\det(\Sigma(t_k)))^{-1/2} \\ &\times \exp\left(-\frac{1}{2}(x - \mu(t_k))^T \Sigma(t_k)^{-1}(x - \mu(t_k))\right). \end{aligned} \quad (2)$$

III. Application 1 - multivariate regression



NN-MaxLike approach

- We approximate the parameter function $\Theta = [\mu, \Sigma]$ via the neural network $\Theta(w) = [\mu(w), \Sigma(w)]$.
- The loss function has the following form

$$\mathcal{L}(w) = \frac{1}{2} \sum_{k=1}^n \left[\ln \left[(2\pi)^d \cdot \det(\Sigma(t_k, w)) \right] + (x_k - \mu(t_k, w))^T \cdot \Sigma^{-1}(t_k, w) \cdot (x_k - \mu(t_k, w)) \right],$$

III. Application 1 - multivariate regression



Step 1. Synthetic data generation - all parameters time-dependent.

- We generate data according to the distribution:

$$X(t_k) = \mu(t_k) + \varepsilon(t_k), \quad k = 0, 1, 2, \dots, n \quad (3)$$

where $\mu(t_k) = [\mu_1(t_k), \mu_2(t_k)]$, $\varepsilon(t_k)$ are independent random variables with the distribution $\mathcal{N}(0, \Sigma(t_k))$, and $t_k = k \frac{2\pi}{n}$, $n = 3000$.

- $X_1 = 0.5 + \sin(t)$, $X_2 = \cos(t)$, $\sigma_1 = 0.1$, $\sigma_2 = 0.15$, $\varrho = 0.5$
- The covariance matrix is as follows

$$\Sigma(t) = \begin{bmatrix} \sigma_1^2 |\mu_1(t)| & \varrho \sigma_1 \sigma_2 |\mu_1(t)| \cdot |\mu_2(t)| \\ \varrho \sigma_1 \sigma_2 |\mu_1(t)| \cdot |\mu_2(t)| & \sigma_2^2 |\mu_2(t)| \end{bmatrix}$$

- $\sigma_1(t) = \sigma_1 \sqrt{\mu_1(t)}$, $\sigma_2(t) = \sigma_2 \sqrt{\mu_2(t)}$, $\varrho(t) = \varrho \sqrt{|\mu_1(t)|} \sqrt{|\mu_2(t)|}$ with $\varrho = 0.5$.
- So the modelled parameter

$$\Theta(t) = [\mu_1(t), \mu_2(t), \sigma_1(t), \sigma_2(t), \varrho(t)].$$

III. Application 1 - multivariate regression



Step 2. Creating and training the neural network model.

Custom loss function - the accurate formula or $d = 2$:

$$\begin{aligned}\mathcal{L}(w) = & \sum_{i=1}^n \left[\ln \left(2\pi\sigma_1(t_i, w)\sigma_1(t_i, w)\sqrt{1-\varrho^2(t_i, w)} \right) \right. \\ & + \frac{1}{2(1-\varrho^2(t_i, w))} \left(\left(\frac{x_i^{(1)} - \mu_1(t_i, w)}{\sigma_1(t_i, w)} \right)^2 + \left(\frac{x_i^{(2)} - \mu_2(t_i, w)}{\sigma_2(t_i, w)} \right)^2 \right. \\ & \left. \left. - 2\varrho(t_i, w) \frac{(x_i^{(1)} - \mu_1(t_i, w))(x_i^{(2)} - \mu_2(t_i, w))}{\sigma_1(t_i, w)\sigma_2(t_i, w)} \right) \right].\end{aligned}$$

III. Application 1 - multivariate regression

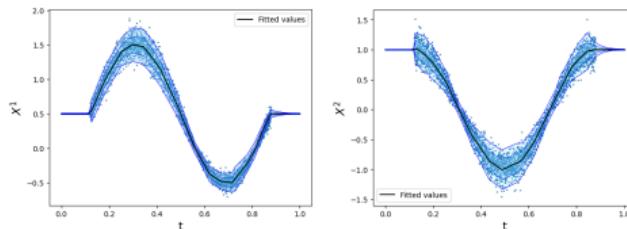


Figure: Fitted values of μ_1 and μ_2 for X_1 and X_2 , respectively, together with 68% and 95% prediction intervals

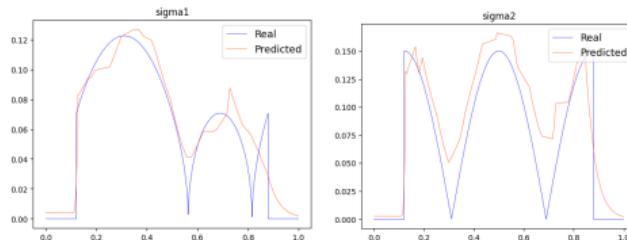


Figure: Parameters reproduced by neural network

IV. Application 2 - SDEs parameters estimation

- SDEs are used for multiple modeling applications - finances, energy consumption, etc.
- We typically don't have access to the SDEs coefficients but just a single observed trajectory
- The typical modeling of real phenomena has two crucial parts - one is choosing the right model, and the second is estimating its parameters (calibration).
- Proposed approach enables us to approximate the underlying SDE based on the single trajectory by using deep learning framework. We only need to assume the type of the SDE, e.g. Black - Scholes.

Calibration of SDEs

$(\Omega, \Sigma, (\Sigma_t)_{t \geq 0}, \mathbb{P}), T > 0, (W(t))_{t \geq 0},$

$$\begin{cases} dX(t) = a(t, X(t), \Theta(t))dt + b(t, X(t), \Theta(t))dW(t), & t \in [0, T], \\ X(0) = x_0 \end{cases} \quad (4)$$

Aim

Having discrete observations x_0, x_1, \dots, x_n of X at $t_i = iT/n$,
 $i = 0, 1, \dots, n$, estimate the parameter function $\Theta : [0, T] \rightarrow \mathbb{R}^s$.

IV. Application 2 - SDEs parameters estimation

Euler scheme: $h = T/n$, $t_k = kh$, $\Delta W_k = W(t_{k+1}) - W(t_k)$

$$X^E(t_0) = x_0$$

$$X^E(t_{k+1}) = X^E(t_k) + a(t_k, X^E(t_k), \theta(t_k))h + b(t_k, X^E(t_k), \theta(t_k))\Delta W_k,$$

$$k = 0, 1, \dots, n-1$$

$(X(t_k))_{k=0,1,\dots,n}$, $(X^E(t_k))_{k=0,1,\dots,n}$ - discrete Markov processes

IV. Application 2 - SDEs parameters estimation

- For $k = 0, 1, \dots, n-1$ we have the transition density for Euler scheme:

$$f_{X^E(t_{k+1})|X^E(t_k)}(x|x_k, \Theta) = (2\pi)^{-d/2} \det(h(b \cdot b^T)(t_k, x_k, \Theta(t_k)))^{-\frac{1}{2}} \\ \times \exp\left[-\frac{1}{2h}(x-x_k-a(t_k, x_k, \Theta(t_k))h)^T(b \cdot b^T)^{-1}(t_k, x_k, \Theta(t_k))(x-x_k-a(t_k, x_k, \Theta(t_k))h)\right]$$

for $x = [x^1, \dots, x^d] \in \mathbb{R}^d$, $x_k = [x_k^1, \dots, x_k^d] \in \mathbb{R}^d$.

- Quasi-MaxLike:

$$\mathcal{L}(\Theta) = - \sum_{k=0}^{n-1} \ln \left[f_{X^E(t_{k+1})|X^E(t_k)}(x_{k+1}|x_k, \Theta) \right] = \\ \frac{1}{2} \sum_{k=0}^{n-1} \left[\ln \left[(2\pi)^d \det(h(b \cdot b^T)(t_k, x_k, \Theta(t_k))) \right] \right. \\ \left. + \frac{1}{h} (\Delta x_k - a(t_k, x_k, \Theta(t_k))h)^T(b \cdot b^T)^{-1}(t_k, x_k, \Theta(t_k)) (\Delta x_k - a(t_k, x_k, \Theta(t_k))h) \right].$$

Number of unknown parameters is $O(n)$, so use NN!

IV. Application 2 - SDEs parameters estimation



NN-Quasi-MaxLike approach:

The loss function for the neural network:

$$\begin{aligned}\tilde{\mathcal{L}}(\mathbf{w}) = & \frac{1}{2} \sum_{k=0}^{n-1} \left[\ln \left[(2\pi)^d \det \left(h(b \cdot b^T)(t_k, x_k, \Theta(t_k, \mathbf{w})) \right) \right] \right. \\ & \left. + \frac{1}{h} \left(\Delta x_k - a(t_k, x_k, \Theta(t_k, \mathbf{w}))h \right)^T (b \cdot b^T)^{-1}(t_k, x_k, \Theta(t_k, \mathbf{w})) \left(\Delta x_k - a(t_k, x_k, \Theta(t_k, \mathbf{w}))h \right) \right],\end{aligned}$$

where $\Delta x_k = x_{k+1} - x_k$.

Aim

$$w^* = \operatorname{argmin}_{w \in \mathbb{R}^N} \tilde{\mathcal{L}}(w)$$

IV. Application 2 - SDEs parameters estimation



EXAMPLE - Ornstein-Uhlenbeck proces with constant θ_1 and varying $\theta_2 = \theta_2(t)$

$$dX(t) = -\theta_1 X(t) dt + \theta_2(t) dW(t), \quad (5)$$

where $\theta_1 > 0$ and $\theta_2 : [0, T] \rightarrow [0, +\infty)$, $\Theta = [\theta_1, \theta_2]$.

Transition density:

$$\begin{aligned} f_{X(t_{k+1})|X(t_k)}(x|x_k, \Theta) &= \left(2\pi \cdot \int_{t_k}^{t_{k+1}} \theta_2^2(u) e^{-2\theta_1 \cdot (t_{k+1}-u)} du \right)^{-1/2} \\ &\times \exp \left[-\frac{(x - e^{-\theta_1 \cdot h} \cdot x_k)^2}{2 \int_{t_k}^{t_{k+1}} \theta_2^2(u) \cdot e^{-2\theta_1 \cdot (t_{k+1}-u)} du} \right]. \end{aligned} \quad (6)$$

IV. Application 2 - SDEs parameters estimation

$$\begin{cases} X(t, w) = a(t, X(t), \Theta(t))dt + b(t, X(t), \Theta(t))dW(t), & t \in [0, T], \\ X(0, w) = x_0 \end{cases} \quad (7)$$

$$a : [0, T] \times \mathbb{R}^d \times \mathbb{R}^s \rightarrow \mathbb{R}^d$$

$$b : [0, T] \times \mathbb{R}^d \times \mathbb{R}^s \rightarrow \mathbb{R}^{d \times m}$$

$$\Theta : [0, T] \rightarrow \mathbb{R}^s$$

- By $\|\cdot\|$ we mean the Euclidean norm in \mathbb{R}^d or \mathbb{R}^s , or the Frobenius norm in $\mathbb{R}^{d \times m}$ (the meaning is clear from the context).
- For $\varrho \geq 0$ we denote by $B(0, \varrho) = \{x \in \mathbb{R}^d \mid \|x\| \leq \varrho\}$.
- For $H \in \mathbb{R}$ we take $H_+ = \max\{H, 0\}$.

We impose the following Krylov-type conditions.

IV. Application 2 - SDEs parameters estimation

Assumptions

(A0) $x_0 \in \mathbb{R}^d$,

(A1) a is Borel measurable and for all $(t, z) \in [0, T] \times \mathbb{R}^s$ the function $a(t, \cdot, z) : \mathbb{R}^d \rightarrow \mathbb{R}^d$ is continuous,

(A2) there exists $H \in \mathbb{R}$ such that for all $t \in [0, T], x \in \mathbb{R}^d, y \in \mathbb{R}^d, z \in \mathbb{R}^s$ it holds

$$\langle x - y, a(t, x, z) - a(t, y, z) \rangle \leq H(1 + \|z\|)\|x - y\|^2,$$

(A3) for all $\varrho \in [0, +\infty)$ it holds

$$\sup_{(t,x) \in [0,T] \times B(0,\varrho)} \|a(t, x, 0)\| < +\infty,$$

(A4) there exist $\alpha_1 \in (0, 1], q \in [1, +\infty), L \in [0, +\infty)$ such that for all $t \in [0, T], x \in \mathbb{R}^d, z_1 \in \mathbb{R}^s, z_2 \in \mathbb{R}^s$ the following holds

$$\|a(t, x, z_1) - a(t, x, z_2)\| \leq L(1 + \|x\|^q)\|z_1 - z_2\|^{\alpha_1},$$

IV. Application 2 - SDEs parameters estimation

Assumptions

- (B1) b is a Borel measurable function,
(B2) there exists $L \in [0, +\infty)$ such that for all $t \in [0, T], x \in \mathbb{R}^d, y \in \mathbb{R}^d, z \in \mathbb{R}^s$ it holds

$$\|b(t, x, z) - b(t, y, z)\| \leq L(1 + \|z\|) \cdot \|x - y\|,$$

(B3)

$$\sup_{0 \leq t \leq T} \|b(t, 0, 0)\| < +\infty,$$

- (B4) there exist $\alpha_2 \in (0, 1], L \in [0, +\infty)$ such that $t \in [0, T], x \in \mathbb{R}^d, z_1 \in \mathbb{R}^s, z_2 \in \mathbb{R}^s$

$$\|b(t, x, z_1) - b(t, x, z_2)\| \leq L(1 + \|x\|) \cdot \|z_1 - z_2\|^{\alpha_2},$$

and finally we assume that

- (T1) Θ is Borel measurable and

$$\sup_{0 \leq t \leq T} \|\Theta(t)\| < +\infty.$$

IV. Application 2 - SDEs parameters estimation

Theorem 1.

Let us assume that $\Theta_i : [0, T] \rightarrow \mathbb{R}^s$, $i = 1, 2$, satisfy the assumption (T1),
 $a : [0, T] \times \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}^d$ satisfies the assumptions (A1)-(A4), and
 $b : [0, T] \times \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}^{d \times m}$ satisfies the assumptions (B1)-(B4). Then

$$\sup_{0 \leq t \leq T} (\mathbb{E} \|X_1(t) - X_2(t)\|^2)^{1/2} \leq C (\|\Theta_1 - \Theta_2\|_\infty^{\alpha_1} + \|\Theta_1 - \Theta_2\|_\infty^{\alpha_2}), \quad (8)$$

where $X_i = X(x_0, a, b, \Theta_i)$, $i = 1, 2$, is the unique strong solution of

$$\begin{cases} dX_i(t) = a(t, X_i(t), \Theta_i(t)) dt + b(t, X_i(t), \Theta_i(t)) dW(t), & t \in [0, T], \\ X_i(0) = x_0, \end{cases} \quad (9)$$

and $C = \max\{C_1, C_2\}$.

We have the following result on existence and uniqueness of strong solution to (7).

Fact 1

Under the assumptions (A0)-(A4), (B1)-(B4), (T1) the SDE (7) has unique strong solution $X = X(x_0, a, b, \Theta)$, such that for all $p \in [2, +\infty)$

$$\mathbb{E} \left(\sup_{0 \leq t \leq T} \|X(t)\|^p \right) < +\infty. \quad (10)$$

IV. Application 2 - SDEs parameters estimation



We stress that C in Theorem 1. does not depend on X_1 . Therefore, if $\|\Theta_1 - \Theta_2\|_\infty \rightarrow 0$ (i.e. $\Theta_1 \rightarrow \Theta_2$ uniformly on $[0, T]$) then

$$\sup_{0 \leq t \leq T} \left(\mathbb{E} \|X(x_0, a, b, \Theta_1)(t) - X(x_0, a, b, \Theta_2)(t)\|^2 \right)^{1/2} \rightarrow 0, \quad (11)$$

and, in particular, $X(x_0, a, b, \Theta_1)(T) \rightarrow X(x_0, a, b, \Theta_2)(T)$ in law.

IV. Application 2 - SDEs parameters estimation

Lemma 2

(Gronwall's lemma with quadratic term, [5]) Let $\varphi : [0, T] \rightarrow [0, +\infty)$ be a bounded Borel measurable function and let there exist $a, \alpha, \beta \in [0, +\infty)$ such that for all $t \in [0, T]$

$$\varphi^2(t) \leq a + 2\alpha \int_0^t \varphi(s) \, ds + 2\beta \int_0^t \varphi^2(s) \, ds. \quad (12)$$

Then for all $t \in [0, T]$

$$\varphi(t) \leq \sqrt{a} \cdot e^{\beta t} + \alpha \cdot \frac{e^{\beta t} - 1}{\beta}. \quad (13)$$

IV. Application 2 - SDEs parameters estimation



PROOF

Let us denote by $\tilde{a}_i(t, x) = a(t, x, \Theta_i(t))$, $\tilde{b}_i(t, x) = b(t, x, \Theta_i(t))$, $i = 1, 2$. Then for all $t \in [0, T]$

$$Z(t) = X_1(t) - X_2(t) = \int_0^t F(s) ds + \int_0^t G(s) dW(s), \quad (14)$$

where $F(s) = \tilde{a}_1(s, X_1(s)) - \tilde{a}_2(s, X_2(s))$, $G(s) = \tilde{b}_1(s, X_1(s)) - \tilde{b}_2(s, X_2(s))$.

From the Itô formula (version from [5], Corollary 2.18) we get that

$$\|Z(t)\|^2 = \int_0^t \left(2\langle Z(s), F(s) \rangle + \|G(s)\|^2 \right) ds + 2 \int_0^t \langle Z(s), G(s) dW(s) \rangle, \quad (15)$$

where

$$\int_0^t \langle Z(s), G(s) dW(s) \rangle = \sum_{k=1}^d \sum_{j=1}^m \int_0^t Z^k(s) \cdot G^{kj}(s) dW^j(s). \quad (16)$$

IV. Application 2 - SDEs parameters estimation



We can show that

$$\begin{aligned} \int_0^T \mathbb{E} [\|Z(t)\|^2 \cdot \|G(t)\|^2] dt &\leq 2TL^2(1 + \|\Theta_1\|_\infty^2) \cdot \mathbb{E}[\|Z\|_\infty^4] \\ &+ 4TL^2(\|\Theta_1\|_\infty + \|\Theta_2\|_\infty)^{2\alpha_2} \cdot \left(\mathbb{E}[\|Z\|_\infty^2] + \mathbb{E}[\|X_2\|_\infty^2 \cdot \|Z\|_\infty^2] \right) < +\infty, \end{aligned}$$

using:

- Assumptions (B2), (B4)
- Auxiliary inequality (based on Fact 1 and Hölder's inequality)

$$\mathbb{E}[\|X_2\|_\infty^2 \cdot \|Z\|_\infty^2] \leq \left(\mathbb{E}[\|X_2\|_\infty^4] \right)^{1/2} \cdot \left(\mathbb{E}[\|Z\|_\infty^4] \right)^{1/2} < +\infty. \quad (17)$$

IV. Application 2 - SDEs parameters estimation

Hence, for all $t \in [0, T]$ $\mathbb{E} \int_0^t \langle Z(s), G(s) dW(s) \rangle = 0$ and thus

$$\mathbb{E}\|Z(t)\|^2 = 2\mathbb{E} \int_0^t \langle Z(s), F(s) \rangle ds + \int_0^t \mathbb{E}\|G(s)\|^2 ds. \quad (18)$$

For all $t \in [0, T]$

$$\|G(t)\| \leq L(1 + \|\Theta_1\|_\infty) \cdot \|Z(t)\| + L(1 + \|X_2\|_\infty) \cdot \|\Theta_1 - \Theta_2\|_\infty^{\alpha_2}, \quad (19)$$

and hence

$$\begin{aligned} \int_0^t \mathbb{E}\|G(s)\|^2 ds &\leq 2L^2(1 + \|\Theta_1\|_\infty)^2 \cdot \int_0^t \mathbb{E}\|Z(s)\|^2 ds \\ &+ 4L^2(1 + \mathbb{E}[\|X_2\|_\infty^2]) \cdot \|\Theta_1 - \Theta_2\|_\infty^{2\alpha_2}. \end{aligned} \quad (20)$$

IV. Application 2 - SDEs parameters estimation

Moreover, for all $t \in [0, T]$, by the assumptions (A2), (A4) and Cauchy-Schwarz inequality

$$\begin{aligned}\langle Z(t), F(t) \rangle &= \langle X_1(t) - X_2(t), a(t, X_1(t), \Theta_1(t)) - a(t, X_2(t), \Theta_2(t)) \rangle \\ &= \langle X_1(t) - X_2(t), a(t, X_1(t), \Theta_1(t)) - a(t, X_2(t), \Theta_1(t)) \rangle \\ &\quad + \langle X_1(t) - X_2(t), a(t, X_2(t), \Theta_1(t)) - a(t, X_2(t), \Theta_2(t)) \rangle \\ &\leq H(1 + \|\Theta_2(t)\|) \cdot \|X_1(t) - X_2(t)\|^2 \\ &\quad + |\langle X_1(t) - X_2(t), a(t, X_2(t), \Theta_1(t)) - a(t, X_2(t), \Theta_2(t)) \rangle| \\ &\leq H_+(1 + \|\Theta_2\|_\infty) \cdot \|Z(t)\|^2 + \|Z(t)\| \cdot \|a(t, X_2(t), \Theta_1(t)) - a(t, X_2(t), \Theta_2(t))\| \\ &\leq H_+(1 + \|\Theta_2\|_\infty) \cdot \|Z(t)\|^2 + L(1 + \|X_2\|_\infty^q) \cdot \|\Theta_1 - \Theta_2\|_\infty^{\alpha_1} \cdot \|Z(t)\|. \quad (21)\end{aligned}$$

IV. Application 2 - SDEs parameters estimation

Therefore, by the Hölder inequality we have for all $t \in [0, T]$

$$\begin{aligned} \mathbb{E} \int_0^t \langle Z(s), F(s) \rangle ds &\leq H_+ (1 + \|\Theta_2\|_\infty) \cdot \int_0^t \mathbb{E} \|Z(s)\|^2 ds \\ &\quad + L \|\Theta_1 - \Theta_2\|_\infty^{\alpha_1} \cdot \int_0^t \mathbb{E} \left[\|Z(s)\| \cdot (1 + \|X_2\|_\infty^q) \right] ds \\ &\leq H_+ (1 + \|\Theta_2\|_\infty) \cdot \int_0^t \mathbb{E} \|Z(s)\|^2 ds \\ &\quad + L \|\Theta_1 - \Theta_2\|_\infty^{\alpha_1} \cdot \left(\mathbb{E} (1 + \|X_2\|_\infty^q)^2 \right)^{1/2} \cdot \int_0^t \left(\mathbb{E} \|Z(s)\|^2 \right)^{1/2} d(s) \end{aligned}$$

IV. Application 2 - SDEs parameters estimation



Combining (20), (18), (22) we get for all $t \in [0, T]$ that

$$\varphi^2(t) \leq a + 2\alpha \int_0^t \varphi(s) \, ds + 2\beta \int_0^t \varphi^2(s) \, ds,$$

where $\varphi(t) = (\mathbb{E}\|Z(t)\|^2)^{1/2}$ is bounded Borel measurable function while

$$a = 4TL^2(1 + \mathbb{E}[\|X_2\|_\infty^2]) \cdot \|\Theta_1 - \Theta_2\|_\infty^{2\alpha_2} \quad (23)$$

$$\alpha = L\|\Theta_1 - \Theta_2\|_\infty^{\alpha_1} \cdot (\mathbb{E}(1 + \|X_2\|_\infty^q)^2)^{1/2} \quad (24)$$

$$\beta = H_+(1 + \|\Theta_2\|_\infty) + L^2(1 + \|\Theta_1\|_\infty)^2. \quad (25)$$

Applying Lemma 2 (Gronwall with quadratic term) we get the thesis. \square

IV. Application 2 - SDEs parameters estimation



AGH

Assumptions - discontinuous drift

In the following scalar case ($m = d = 1$) with additive noise

$$\begin{cases} dX(t) = a(t, X(t)) dt + \Theta(t) dW(t), & t \in [0, T], \\ X(0) = x_0, \end{cases} \quad (26)$$

we consider the following *Zvonkin-type* for $a : [0, T] \times \mathbb{R} \rightarrow \mathbb{R}$ and $\Theta : [0, T] \rightarrow \mathbb{R}$:

- (C0) $x_0 \in \mathbb{R}$,
- (C1) a is Borel measurable,
- (C2) there exists $H \in \mathbb{R}$ such that for all $t \in [0, T]$, $x, y \in \mathbb{R}$ it holds

$$(x - y)(a(t, x) - a(t, y)) \leq H|x - y|^2, \quad (27)$$

- (C3) there exists $D \in [0, +\infty)$ such that for all $t \in [0, T]$, $x \in \mathbb{R}$ it holds

$$|a(t, x)| \leq D, \quad (28)$$

- (D1) Θ is Borel measurable,
- (D2) there exist $d_0, d_1 \in (0, +\infty)$ such that for all $t \in [0, T]$

IV. Application 2 - SDEs parameters estimation

Theorem 2 - discontinuous drift

Let us assume that $\Theta_i : [0, T] \rightarrow \mathbb{R}$, $i = 1, 2$, satisfy the assumptions $(D1), (D2)$, and $a : [0, T] \times \mathbb{R} \rightarrow \mathbb{R}$ satisfies the assumptions $(C1)-(C3)$. Then

$$\sup_{0 \leq t \leq T} \left(\mathbb{E} \|X_1(t) - X_2(t)\|^2 \right)^{1/2} \leq e^{TH_+} \|\Theta_1 - \Theta_2\|_{L^2[0, T]}, \quad (30)$$

where $X_i = X(x_0, a, \Theta_i)$, $i = 1, 2$, is the unique strong solution of

$$\begin{cases} dX_i(t) = a(t, X_i(t)) dt + \Theta_i(t) dW(t), & t \in [0, T], \\ X_i(0) = x_0. \end{cases} \quad (31)$$

IV. Application 2 - SDEs parameters estimation

Example 1 - Ornstein-Uhlenbeck process

$$dX(t) = \kappa(\mu - X(t))dt + \sigma(t)dW(t) \quad (32)$$

$\kappa = 2$, $\mu = 0.5$, $\sigma(t) = 2t + 0.4 + 1.5 \cdot \sin(4t)$.

Modelled parameter: $\theta(t) = \sigma(t)$.

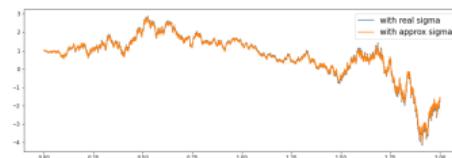
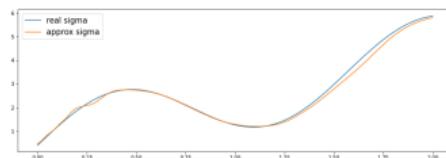


Figure: Parameter $\sigma(t)$ for $t \in [0, T]$ estimated with neural network and trajectories of the process with real and approximated values. Generated by the Euler-Maruyama scheme with the step-size $h = 0.0002$.

IV. Application 2 - SDEs parameters estimation

Example 2 - discontinuous Ornstein-Uhlenbeck process

We consider the following SDE (also known as threshold diffusion)

$$dX(t) = (\mu - \kappa \cdot \text{sign}(X(t))) dt + \sigma(t) dW(t), \quad (33)$$

$$\kappa = 2, \mu = 0.5, \sigma(t) = 2t + 0.4 + 1.5 \cdot \sin(4t).$$

Modelled parameter: $\theta(t) = \sigma(t)$.

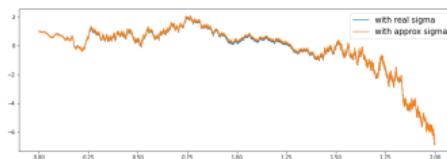
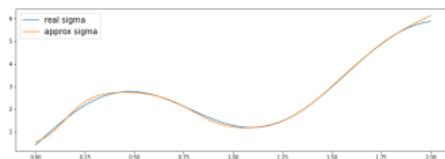


Figure: Parameter $\sigma(t)$ for $t \in [0, T]$ estimated with neural network and trajectories of the process with real and approximated values. Generated by the Euler-Maruyama scheme with the step-size $h = 0.0002$.

IV. Application 2 - SDEs parameters estimation

Example 3 - SDE with nonlinear coefficients

$$dX(t) = \kappa \cdot \cos(X(t))dt + \left((\sin(X(t)) + 1.5) \cdot \sigma(t) + 2 \right) dW(t), \quad (34)$$

$\kappa = 0.4$ and $\sigma(t) = 2 \cdot \sin(2\pi t) + t$.

Modelled parameter: $\theta(t) = \sigma(t)$, κ is known to the algorithm.

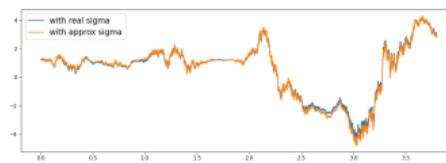
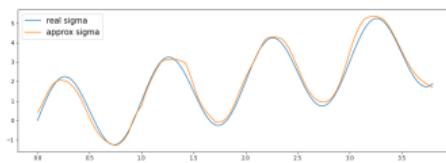


Figure: Parameter $\sigma(t)$ for $t \in [0, T]$ estimated with neural network and trajectories of the process with real and approximated values. Generated by the Euler-Maruyama scheme with the step-size $h = 0.00038$.

IV. Application 2 - SDEs parameters estimation

Evaluation of the results - histograms and qq-plots

To evaluate our methodology, we need to compare the trajectories generated with two sets of coefficients - real and approximated by the neural network. Proposed approach is to simulate multiple trajectories ($N = 1000$) for each option and compare both distributions of value $X(T)$. Our aim is to verify that the distributions are similar.

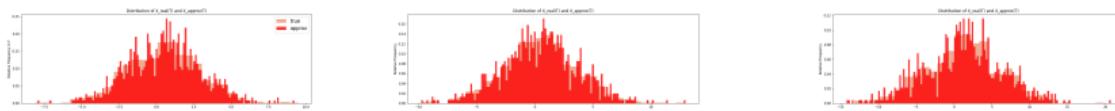


Figure: Histogram of values for real $X(T)$ and its approximation with computed $\sigma(t)$.

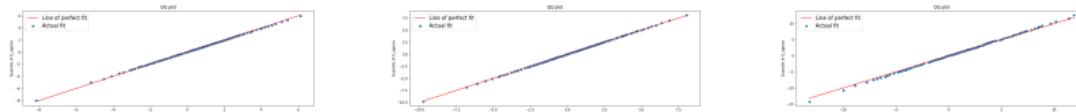


Figure: QQ plots - comparison of both distributions.

IV. Application 2 - SDEs parameters estimation

	Example 1		Example 2		Example 3	
	X real	X approx	X real	X approx	X real	X approx
empirical mean	0.515	0.515	0.495	0.497	1.243	1.256
empirical std dev	2.456	2.388	3.168	3.202	4.835	5.077

Table: Empirical distribution values for each example - mean and standard deviation

IV. Application 2 - SDEs parameters estimation

Kolmogorov-Smirnov test with the null hypothesis that both the distributions P , Q are identical based on statistic $D_{n,m}$

$$D_{n,m} = \max |F_P(x) - F_Q(x)| \quad (35)$$

$$n = m = 1000$$

F_P is CDF of P and F_Q CDF of Q

	Example 1	Example 2	Example 3
KS statistics	0.015	0.009	0.028
p-value of KS test	0.99987	0.99999	0.82821

Table: Comparison of distributions metrics for all the examples

$$D_{n,m} = \max |F_P(x) - F_Q(x)| \quad (36)$$

No reason to reject null hypothesis for E1 and E2 ($\alpha = 0.05$).

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Thank you for your attention!

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